

Soc 63993, Homework #5 Answer Key: Model Mis-Specification/Equality Constraints/Group Comparisons

Richard Williams, University of Notre Dame, <https://academicweb.nd.edu/~rwilliam/>

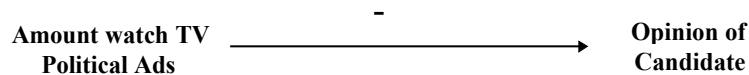
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1. **Model mis-specification.** A campaign manager has found that the amount of time spent watching TV political ads is negatively correlated with favorable opinion of her candidate. Two models have been proposed to explain this relationship:

(i)



(ii)



A. Suppose that model (i) is correct. What harm will result from estimating model (ii) and relying on the results? If appropriate, discuss such things as biased coefficients, inflated standard errors, and misguided policy decisions (particularly with regards to the use of TV advertising). Similarly, discuss the harm that will result if Model (ii) is correct and model (i) is mistakenly estimated and relied upon.

If model (ii) is estimated, the data will seem to support it even if Model (i) is correct. Model two predicts a negative effect of TV on opinion. Since the correlation between TV and opinion is negative, the bivariate regression coefficient will indeed be negative.

However, as the “true” model (i) shows, the effect of TV on opinion is actually positive. The negative correlation between TV and opinion arises from the fact that they share a common cause, Education. Better educated people are less likely to watch TV ads and more likely to like the candidate. Hence, those who watch TV ads tend to be disproportionately composed of the lesser-educated who are less likely to like the candidate. However, they would like the candidate even less if they didn’t watch the TV ads. Put another way, suppressor effects are present.

From a policy standpoint, this could lead to a grave mistake. The campaign could mistakenly conclude that it should turn away from TV advertising, when that advertising is actually helpful.

Probably less harm is done if model (ii) is correct but model (i) is estimated instead. The expected effect of Education on Opinion is zero, and the expected effect of TV on Opinion is negative. That is, adding extraneous variables to the model does not bias coefficients. Hence, when model (i) is estimated, the campaign will hopefully discover that the data don’t support it (whereas in the previous case the data did seem to support the model, even though it was wrong). However, adding extraneous variable does tend to increase standard errors and make estimates less precise. Hence, there is a greater risk that the campaign will conclude that TV does not have

a significant effect (i.e. it is “neutral”) when in fact the TV ads are harmful. For that matter, because the estimates are less precise, there may even be a small chance that the estimated effect of TV winds up being positive, as Model (i) suggests.

B. Model (i) is estimated, yielding the following results. Based on this information, determine what the regression coefficient would be for model (ii). Compute the regression coefficient using both the formula for omitted variable bias and the formula for the slope coefficient in a bivariate regression.

```
. sum
```

Variable	Obs	Mean	Std. Dev.	Min	Max
opinion	200	79	9.4	57.15274	99.79181
educ	200	14	2.7	6.597328	20.61872
tv	200	15	5.6	-.8872261	34.85936

```
. corr
```

```
(obs=200)
```

	opinion	educ	tv
opinion	1.0000		
educ	0.3500	1.0000	
tv	-0.2200	-0.9000	1.0000

```
. corr, cov
```

```
(obs=200)
```

	opinion	educ	tv
opinion	88.36		
educ	8.883	7.29	
tv	-11.5808	-13.608	31.36

```
. reg opinion educ tv, beta
```

Source	SS	df	MS	Number of obs =	200
Model	2989.21851	2	1494.60926	F(2, 197) =	20.17
Residual	14594.421	197	74.0833553	Prob > F =	0.0000
Total	17583.6395	199	88.3599975	R-squared =	0.1700
				Adj R-squared =	0.1616
				Root MSE =	8.6072

	opinion	Coef.	Std. Err.	t	P> t	Beta
	educ	2.785185	.5184337	5.37	0.000	.8
	tv	.8392856	.2499591	3.36	0.001	.5
	_cons	27.41812	10.77459	2.54	0.012	.

We are asked to compute the coefficients for the incorrectly-specified bivariate regression. I’ll do this for both TV and Education as the IVs.

Opinion regressed on TV only

$$b = \frac{s_{TV, Opinion}}{s_{TV}^2} = \frac{-11.581}{31.360} = -.369$$

$$b_{TV}^* = b_{TV} + b_{Educ} \frac{Cov(TV, EDUC)}{V(TV)}$$

$$= .839286 + 2.785185 \frac{-13.608}{31.360} = -.369$$

(Bonus) Opinion regressed on education only

$$b = \frac{s_{Educ, Opinion}}{s_{Educ}^2} = \frac{8.883}{7.29} = 1.219$$

$$b_{Educ}^* = b_{Educ} + b_{TV} \frac{Cov(TV, EDUC)}{V(Educ)}$$

$$= 2.785185 + .839286 \frac{-13.608}{7.290} = 1.219$$

To confirm – note that we are given the means, correlations and standard deviations, so we can use the `corr2data` command to create a pseudo-replication of the data.

```
. matrix input means = (79\14\15)
. matrix input sds = (9.4\2.7\5.6)
. matrix input corr = (1, .35, -.22\-.35, 1, -.90\-.22, -.90, 1)
. corr2data opinion educ tv, n(200) means(means) sds(sds) corr(corr)
. reg opinion educ tv, beta
```

Source	SS	df	MS	Number of obs =	200
Model	2989.2186	2	1494.6093	F(2, 197) =	20.17
Residual	14594.4214	197	74.0833575	Prob > F =	0.0000
				R-squared =	0.1700
				Adj R-squared =	0.1616
Total	17583.64	199	88.3600001	Root MSE =	8.6072

opinion	Coef.	Std. Err.	t	P> t	Beta
educ	2.785185	.5184337	5.37	0.000	.8
tv	.8392857	.2499591	3.36	0.001	.5
_cons	27.41812	10.77459	2.54	0.012	.

```
. reg opinion tv, beta
```

Source	SS	df	MS	Number of obs =	200
Model	851.048099	1	851.048099	F(1, 198) =	10.07
Residual	16732.5919	198	84.50804	Prob > F =	0.0017
				R-squared =	0.0484
				Adj R-squared =	0.0436
Total	17583.64	199	88.3600001	Root MSE =	9.1928

opinion	Coef.	Std. Err.	t	P> t	Beta
tv	-.3692857	.1163682	-3.17	0.002	-.22
_cons	84.53929	1.862631	45.39	0.000	.

```
. reg opinion educ, beta
```

Source	SS	df	MS		Number of obs =	200
Model	2153.99576	1	2153.99576		F(1, 198) =	27.64
Residual	15429.6443	198	77.9274963		Prob > F =	0.0000
					R-squared =	0.1225
					Adj R-squared =	0.1181
Total	17583.64	199	88.3600001		Root MSE =	8.8277

opinion	Coef.	Std. Err.	t	P> t	Beta
educ	1.218518	.2317688	5.26	0.000	.35
_cons	61.94074	3.304259	18.75	0.000	.

C. Based on these results, which model do you think is most plausible? Why?

Model (i) gets a clear edge. All coefficients are in the predicted direction, and all effects are statistically significant.

D. The campaign manager is concerned by the large correlation between educ and tv. Suppose the manager decided to “solve” the problem of multicollinearity by excluding education from the model. What would be the consequence of that decision? Do you think this would be a good idea in this case?

It would be a terrible mistake if you decided to “solve” the problem of multicollinearity by excluding education from the model. As noted above, this serious mis-specification would lead to very erroneous conclusions concerning TV ads. Further, even with this high correlation, effects are statistically significant. Stick with Model (i).

Incidentally, keep in mind that omitted variable bias can cause the magnitude of the coefficients for the remaining variables to be inflated either upwards or downwards. In this case, omitting education would cause the effect of TV to go down so much that the estimated effect actually switches from being positive to negative. This is because there are suppressor effects present in this example: TV and Education both positively affect opinion, but they are negatively correlated with each other.

2. Equality constraints. From the course web page, download gender.dta. This is yet another modified version of our income/education/job experience example. The sample now consists of 225 men and 275 women. Regress income on education and job experience. Test the following hypotheses:

$$H_0: \beta_{\text{Educ}} = \beta_{\text{Jobexp}}$$

$$H_A: \beta_{\text{Educ}} \neq \beta_{\text{Jobexp}}$$

Perform a Wald test, an incremental F test, and a likelihood ratio chi-square test. The results should all be identical or nearly identical.

(i) Wald test:

```
. use https://academicweb.nd.edu/~rwilliam/xsoc63993/statafiles/gender.dta, clear
. * Unconstrained model
. reg income educ jobexp
```

Source	SS	df	MS	Number of obs =	500
Model	22352.7545	2	11176.3773	F(2, 497) =	239.86
Residual	23157.8824	497	46.5953368	Prob > F =	0.0000
				R-squared =	0.4912
				Adj R-squared =	0.4891
Total	45510.6369	499	91.2036811	Root MSE =	6.8261

income	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
educ	1.309229	.0838474	15.61	0.000	1.14449 1.473968
jobexp	.8533107	.0670888	12.72	0.000	.7214982 .9851233
_cons	-1.076636	1.205717	-0.89	0.372	-3.445568 1.292295

```
. test educ = jobexp
```

```
( 1) educ - jobexp = 0
```

```
F( 1, 497) = 15.63
Prob > F = 0.0001
```

To confirm that Stata got it right:

```
. vce
```

	educ	jobexp	_cons
educ	.00703		
jobexp	-.000883	.004501	
_cons	-.065025	-.049566	1.45375

$$F_{1, N-K-1} = \left(\frac{(b_{Educ} - b_{Jobexp})}{\sqrt{s^2_{b_{Educ}} + s^2_{b_{Jobexp}} - 2s_{b_{Educ}, b_{Jobexp}}}} \right)^2 = \left(\frac{(1.309229 - .8533107)}{\sqrt{.00703 + .004501 - 2 * -.000883}} \right)^2$$

$$= \left(\frac{.4559183}{.115312619} \right)^2 = 3.95375897^2 = 15.63$$

(ii) Incremental F test:

```
. * Unconstrained model
. reg income educ jobexp
```

Source	SS	df	MS	Number of obs	=	500
Model	22352.7545	2	11176.3773	F(2, 497)	=	239.86
Residual	23157.8824	497	46.5953368	Prob > F	=	0.0000
				R-squared	=	0.4912
				Adj R-squared	=	0.4891
Total	45510.6369	499	91.2036811	Root MSE	=	6.8261

income	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
educ	1.309229	.0838474	15.61	0.000	1.14449 1.473968
jobexp	.8533107	.0670888	12.72	0.000	.7214982 .9851233
_cons	-1.076636	1.205717	-0.89	0.372	-3.445568 1.292295

```
. est store unconstrained
. * Constrained model
. gen jobed = educ + jobexp
. reg income jobed
```

Source	SS	df	MS	Number of obs	=	500
Model	21624.34	1	21624.34	F(1, 498)	=	450.84
Residual	23886.2969	498	47.9644516	Prob > F	=	0.0000
				R-squared	=	0.4751
				Adj R-squared	=	0.4741
Total	45510.6369	499	91.2036811	Root MSE	=	6.9256

income	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
jobed	1.037904	.0488816	21.23	0.000	.9418644 1.133944
_cons	-.5465906	1.215718	-0.45	0.653	-2.935159 1.841978

```
. est store constrained
. * Use Buis's -fctest- command
. fctest constrained unconstrained
Assumption: constrained nested in unconstrained
```

```
F( 1, 497) = 15.63
prob > F = 0.0001
```

If you prefer to do things the hard way – From the unconstrained model, we get

$$SSE_u = 23157.8824, R_u^2 = .4912, N = 500, K = 2.$$

From the constrained model, we get

$$SSE_c = 23886.2969, R_c^2 = .4751, J = 1.$$

Using the incremental F test, we get

$$F_{1,N-K-1} = \frac{(SSE_c - SSE_u) * (N - K - 1)}{SSE_u * 1} = \frac{(R_u^2 - R_c^2) * (N - K - 1)}{(1 - R_u^2) * 1}$$

$$= \frac{(23886.30 - 23157.88) * 497}{23157.88} = \frac{(.4912 - .4751) * 497}{1 - .49115} = 15.63$$

(iii) Likelihood ratio chi square test:

```
. lrtest constrained unconstrained, stats
```

```
Likelihood-ratio test                                LR chi2(1)  =    15.48
(Assumption: constrained nested in unconstrained)    Prob > chi2 =    0.0001
```

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll (null)	ll (model)	df	AIC	BIC
constrained	500	-1837.243	-1676.082	2	3356.165	3364.594
unconstrained	500	-1837.243	-1668.34	3	3342.68	3355.324

Note: N=Obs used in calculating BIC; see [R] BIC note

The test statistics are all highly significant. It is very unlikely that the effects of education and job experience are equal.

3. Group comparisons. Using the same data as in problem 2, do the following:

(a) Do T-tests of whether the means of men and women significantly differ on education, job experience, and income. If using Stata, use commands such as

```
. ttest educ, by(female)
```

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
male	225	11.22222	.298438	4.47657	10.63412	11.81033
female	275	10.63636	.1733252	2.874273	10.29515	10.97758
combined	500	10.9	.1650287	3.690154	10.57576	11.22424
diff		.5858586	.3310136		-.0644967	1.236214

Degrees of freedom: 498

Ho: mean(male) - mean(female) = diff = 0

Ha: diff < 0	Ha: diff != 0	Ha: diff > 0
t = 1.7699	t = 1.7699	t = 1.7699
P < t = 0.9613	P > t = 0.0774	P > t = 0.0387

Men have slightly more education than women do. The difference is significant if you use a 1-tailed test.

```
. ttest jobexp, by(female)
```

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
male	225	14.11111	.3569664	5.354497	13.40767	14.81455
female	275	12.36364	.2249718	3.730735	11.92074	12.80653
combined	500	13.15	.2062525	4.611945	12.74477	13.55523
diff		1.747475	.4075443		.9467565	2.548193

Degrees of freedom: 498

Ho: mean(male) - mean(female) = diff = 0

Ha: diff < 0	Ha: diff != 0	Ha: diff > 0
t = 4.2878	t = 4.2878	t = 4.2878
P < t = 1.0000	P > t = 0.0000	P > t = 0.0000

On average, men have almost 2 more years of job experience than do women. The difference is highly significant.

```
. ttest income, by(female)
```

Two-sample t test with equal variances

Group	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
male	225	27.81111	.76553	11.48295	26.30255	29.31967
female	275	21.63636	.3865032	6.40943	20.87547	22.39726
combined	500	24.415	.4270917	9.550062	23.57588	25.25412
diff		6.174747	.8135835		4.576268	7.773227

Degrees of freedom: 498

Ho: mean(male) - mean(female) = diff = 0

Ha: diff < 0	Ha: diff != 0	Ha: diff > 0
t = 7.5896	t = 7.5896	t = 7.5896
P < t = 1.0000	P > t = 0.0000	P > t = 0.0000

Men make more than \$6,000 a year more than women, and the difference is highly significant.

(b) Test the following. Use a likelihood ratio chi square test. Performing an incremental F test and/or a Wald test using suest is optional.

H₀: Model parameters are the same for both men and women
H_A: Model parameters are not the same for both men and women.

```
. * Constrained model: No Gender differences
. reg income educ jobexp
```

Source	SS	df	MS	Number of obs = 500		
Model	22352.7545	2	11176.3773	F(2, 497)	=	239.86
Residual	23157.8824	497	46.5953368	Prob > F	=	0.0000
				R-squared	=	0.4912
				Adj R-squared	=	0.4891
				Root MSE	=	6.8261
Total	45510.6369	499	91.2036811			

income	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
educ	1.309229	.0838474	15.61	0.000	1.14449	1.473968
jobexp	.8533107	.0670888	12.72	0.000	.7214982	.9851233
_cons	-1.076636	1.205717	-0.89	0.372	-3.445568	1.292295

Note that the constrained model was the unconstrained model in problem 2. In problem 2, we viewed it as unconstrained because the effects of education and job experience were free to differ. In this problem, we view it as constrained because the coefficients are constrained to be the same for both men and women.

```
. * Unconstrained - Effects differ by gender
. reg income educ jobexp if female == 0
```

Source	SS	df	MS	Number of obs = 225		
Model	19350.4582	2	9675.22912	F(2, 222)	=	210.87
Residual	10185.7638	222	45.8818188	Prob > F	=	0.0000
				R-squared	=	0.6551
				Adj R-squared	=	0.6520
				Root MSE	=	6.7736
Total	29536.222	224	131.858134			

income	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
educ	.8195378	.1070818	7.65	0.000	.6085108	1.030565
jobexp	1.384972	.0895246	15.47	0.000	1.208545	1.561398
_cons	-.9294128	1.49777	-0.62	0.536	-3.88108	2.022254

```
. est store male
```

```
. reg income educ jobexp if female == 1
```

Source	SS	df	MS	Number of obs = 275		
Model	5276.94296	2	2638.47148	F(2, 272)	=	120.03
Residual	5979.19312	272	21.9823276	Prob > F	=	0.0000
				R-squared	=	0.4688
				Adj R-squared	=	0.4649
				Root MSE	=	4.6885
Total	11256.1361	274	41.0807886			

income	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
educ	1.525582	.1004096	15.19	0.000	1.327903	1.723261
jobexp	-.0049199	.0773587	-0.06	0.949	-.1572178	.1473779
_cons	5.470545	1.589722	3.44	0.001	2.340821	8.600269

```
. est store female
```

Likelihood Ratio Test. Doing a Likelihood Ratio Chi Square Test,

```
. lrtest (male female) both
```

Likelihood-ratio test

```
LR chi2(3) = 213.10  
Prob > chi2 = 0.0000
```

Assumption: (both) nested in (male, female)

[Note that the LR Chi Square / Degrees of Freedom = $213.10/3 = 71.03$; compare with the incremental F value calculated below]

This is highly significant, ergo we reject the null and conclude that the coefficients for men and women likely are different. This conclusion is not surprising, since, by looking at the coefficients in the separate male and female models, you can see that the effects appear to be very different. Nonetheless, keep in mind that all we know for sure is that at least one parameter (including possibly the intercept) differs between men and women.

Incremental F Test. If we want to be masochistic and do an incremental F test, from the constrained model we get

$$SSE_c = 23158, N = 500, K = 2.$$

From the regressions for males only and females only we get

$$SSE_{\text{Males}} = 10186, N_{\text{Males}} = 225$$

$$SSE_{\text{Females}} = 5979, N_{\text{Females}} = 275.$$

Hence, by adding up the figures for men and women, for the unconstrained model we get

$$SSE_u = 16165, N_u = 500.$$

Also, note that $J = K + 1 = 3$, i.e. the constrained model estimates 2 betas and 1 intercept, while the unconstrained model estimates 4 betas and 2 intercepts.

Hence, for the incremental F, we get

$$F_{K+1, N_1 + N_2 - 2K - 2} = \frac{(SSE_c - SSE_u) * (N_1 + N_2 - 2K - 2)}{SSE_u * (K + 1)} = \frac{(23158 - 16165) * 494}{16165 * 3} = 71.24$$

[Note too that this is almost identical to LR chi square/3 shown above]

Suest. If we wanted to do this with a Wald chi-square test and the `suest` command,

```
. quietly reg income educ jobexp if female == 0
. est store male
. quietly reg income educ jobexp if female == 1
. est store female
. suest male female
```

Simultaneous results for male, female

Number of obs = 500

	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
male_mean						
educ	.8195378	.1000803	8.19	0.000	.623384	1.015692
jobexp	1.384972	.1096212	12.63	0.000	1.170118	1.599825
_cons	-.9294128	.494266	-1.88	0.060	-1.898156	.0393307
male_lnvar						
_cons	3.826069	.0705412	54.24	0.000	3.687811	3.964327
female_mean						
educ	1.525582	.0930839	16.39	0.000	1.343141	1.708023
jobexp	-.0049199	.0400294	-0.12	0.902	-.0833761	.0735362
_cons	5.470545	1.626955	3.36	0.001	2.281772	8.659318
female_lnvar						
_cons	3.090239	.1010872	30.57	0.000	2.892112	3.288366

```
. test [male_mean = female_mean], constant coef
```

```
( 1) [male_mean]educ - [female_mean]educ = 0
( 2) [male_mean]jobexp - [female_mean]jobexp = 0
( 3) [male_mean]_cons - [female_mean]_cons = 0
```

```
chi2( 3) = 180.32
Prob > chi2 = 0.0000
```

Constrained coefficients

	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
male_mean						
educ	1.579664	.0309782	50.99	0.000	1.518948	1.64038
jobexp	.0646106	.0225819	2.86	0.004	.0203509	.1088703
_cons	3.960761	.3267024	12.12	0.000	3.320436	4.601085
male_lnvar						
_cons	3.813277	.0572104	66.65	0.000	3.701146	3.925407
female_mean						
educ	1.579664	.0309782	50.99	0.000	1.518948	1.64038
jobexp	.0646106	.0225819	2.86	0.004	.0203509	.1088703
_cons	3.960761	.3267024	12.12	0.000	3.320436	4.601085
female_lnvar						
_cons	2.949792	.092498	31.89	0.000	2.768499	3.131085

(c) Based on your results, explain whether men make more than women and if so why. [Note: these are hypothetical data, and the results are a little peculiar in some respects!]

We know from the T-Tests that, on average, men make significantly more money than do women. We also know from the T-Tests that men benefit from having higher levels of education and job experience than do women. The regressions add additional insights as to why differences exist. For men, both education and job experience have significant effects, with job experience actually having a larger effect than education does. For women, on the other hand, job experience has virtually no effect whatsoever; only education is important. Education actually appears to have a larger effect on women than it does men! But this is more than offset by the advantages men have from higher levels of job experience and education and the much greater effect job experience has on men than women. Perhaps women are more likely to be in dead-end jobs where additional experience does not help you to get promoted into higher paying positions.

It is true that, under certain conditions, a woman would be expected to make more than a comparable man, e.g. when $\text{jobexp} = 0$. However, no such person exists in the sample (the lowest value of jobexp is 3), and overall, men have the advantage.

These would be extremely interesting and important findings, if it weren't for the fact that I made these data up.

(d) Suppose there were no gender-related compositional differences, i.e. women had the same levels of education and job experience as men did. If education and job experience continued to have the same effects on women that they do now, how much would the gap in income between men and women be affected?

We are asking a “what if” question. The following analysis addresses this.

```
. tabstat income educ jobexp, by(female) columns(variables)
```

Summary statistics: mean
by categories of: female

female	income	educ	jobexp
male	27.81111	11.22222	14.11111
female	21.63636	10.63636	12.36364
Total	24.415	10.9	13.15

As we saw before, men make \$6174.75 more than women on average.

```
. reg income educ jobexp if female == 1
```

Source	SS	df	MS	Number of obs =	275
Model	5276.94296	2	2638.47148	F(2, 272) =	120.03
Residual	5979.19312	272	21.9823276	Prob > F =	0.0000
Total	11256.1361	274	41.0807886	R-squared =	0.4688
				Adj R-squared =	0.4649
				Root MSE =	4.6885

income	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
educ	1.525582	.1004096	15.19	0.000	1.327903 1.723261
jobexp	-.0049199	.0773587	-0.06	0.949	-.1572178 .1473779
_cons	5.470545	1.589722	3.44	0.001	2.340821 8.600269

```
. margins, at (educ = 11.2222 jobexp = 14.11111)
```

```
Adjusted predictions      Number of obs   =      275
Model VCE      : OLS
```

```
Expression   : Linear prediction, predict()
at           : educ           =    11.2222
              jobexp          =    14.11111
```

		Delta-method				
	Margin	Std. Err.	t	P> t	[95% Conf. Interval]	
_cons	22.52151	.3236064	69.60	0.000	21.88442	23.1586

As the final numbers show, if women had the same levels of education and job experience as men, while the effects of education and job experience on women stayed the same, women would make \$22,521.54 on average, an increase of $\$22,521.54 - \$21,636.36 = \$885.18$. Hence, of the original difference of \$6174.75, $\$885.18 / \$6174.75 = 14.34\%$ is due to compositional factors. The rest is due to differences in effects, in particular the fact that women get virtually no benefit from their years of job experience.

Some alternative approaches that will also work:

Using the predict command,

```
. reg income educ jobexp if female == 1
. predict mcompfcoef if !female
(option xb assumed; fitted values)
(275 missing values generated)
. sum mcompfcoef
```

Variable	Obs	Mean	Std. Dev.	Min	Max
mcompfcoef	225	22.52154	6.82074	8.506949	31.35624

Using the adjust command,

```
. quietly reg income educ jobexp if female == 1
. adjust educ jobexp if female == 0
```

```
-----
Dependent variable: income      Command: regress
Covariates set to mean: educ = 11.222222, jobexp = 14.11111
-----
```

All	xb
	22.5215

```
Key:  xb = Linear Prediction
```

```
. quietly reg income educ jobexp if female == 1
. margins if female == 0, atmeans noesample
```

Adjusted predictions	Number of obs	=	225
Model VCE : OLS			

```
Expression : Linear prediction, predict()
at         : educ           = 11.22222 (mean)
           : jobexp        = 14.11111 (mean)
```

	Delta-method					
	Margin	Std. Err.	t	P> t	[95% Conf. Interval]	
_cons	22.52154	.323607	69.60	0.000	21.88445	23.15863

Using margins with precise values,

```
. sum educ if female == 0, meanonly
. scalar malemeaneduc = r(mean)
. sum jobexp if female == 0, meanonly
. scalar malemeanjobexp = r(mean)
. scalar list
```

```
malemeanjobexp = 14.111111
malemeaneduc = 11.222222
```

```
. quietly reg income educ jobexp if female == 1
. margins, at (educ = `malemeaneduc' jobexp = `malemeanjobexp')
```

Adjusted predictions	Number of obs	=	275
Model VCE : OLS			

```
Expression : Linear prediction, predict()
at          : educ              = 11.22222
              jobexp            = 14.11111
```

	Delta-method					
	Margin	Std. Err.	t	P> t	[95% Conf. Interval]	
_cons	22.52154	.323607	69.60	0.000	21.88445	23.15863