Specification Error: Omitted and Extraneous Variables

Richard Williams, University of Notre Dame, https://academicweb.nd.edu/~rwilliam/ Last revised February 15, 2015

Omitted variable bias. Suppose that the "correct" model is

$$y = \alpha + \beta_1 X_1 + \beta_2 X_2 + \varepsilon$$

If we estimate

$$y = a + b_1 X_1 + b_2 X_2 + e$$

we know that $E(b_1) = \beta_1$ and $E(b_2) = \beta_2$ i.e. the regression coefficients are unbiased estimators of the population parameters.

Suppose, however, the researcher mistakenly believes

$$y = \alpha^* + \beta^*_{1}X_{1} + \varepsilon^*$$

and therefore estimates

$$y = a^* + b^*_1 X_1 + e^*$$

i.e. X2 is mistakenly omitted from the model. How does b₁ (the regression estimate from the correctly specified model) compare to b₁* (the regression estimate from the mis-specified model)? What is $E(b_1^*)$? Is it a biased or unbiased estimator of β_1 ? If biased, how is it biased?

Note that b₁*

$$= \frac{\hat{C}ov(X_1, Y)}{\hat{V}(X_1)}$$

$$= \frac{\hat{C}ov(X_1, a + b_1X_1 + b_2X_2 + e)}{\hat{V}(X_1)}$$
Substitute the from the commodel
$$= \frac{\hat{C}ov(X_1, a) + b_1\hat{C}ov(X_1, X_1) + b_2\hat{C}ov(X_1, X_2) + \hat{C}ov(X_1, e)}{\hat{V}(X_1)}$$
Expectations Cov(a+b,c+c) Cov(a,d) + Cov(a,d) +

Formula for bivariate regression

Substitute the formula for Y from the correctly specified

Expectations rules: Cov(a+b,c+d) = Cov(a,c) +Cov(a,d) + Cov(b,c) + Cov(b,d)

Recall that Cov(variable, constant) = 0. Also, X's are uncorrelated with the residuals.

Simplify expression.

If your eyes glaze over when looking at equations, just make sure you get the conclusion. If X2 has mistakenly been omitted from the model, then, taking expectations, we get

$$E(b_1^*) = \beta_1 + \beta_2 \frac{\sigma_{12}}{\sigma_1^2}$$

Very Important: Hence, b_1^* is a biased estimator of β_1 . Further, this bias will not disappear as sample size gets larger, so the omission of a variable from a model also leads to an inconsistent estimator. In effect, x1 gets credit (or blame) for the effects of the variables that have been omitted from the model.

Note that there are two conditions under which b_1^* will not be biased:

- $\beta_2 = 0$. Of course, if $\beta_2 = 0$, this means that the model is not mis-specified, i.e. X2 does not belong in the model because it has no effect on Y.
- $\sigma_{12} = 0$. That is, if the 2 X's are uncorrelated, then omitting one does not result in biased estimates of the effect of the other.

Example 1. I will construct a data set where b1 = 3, b2 = 2, and x1 and x2 have a correlation of .5. The standard deviation of x1 is 4 and the standard deviation of x2 is 4. We will see what happens if x2 is omitted from the model.

```
. clear all
. matrix input corr = (1,.5,0 \setminus .5,1,0 \setminus 0,0,1)
. matrix input sds = (4\10)
. corr2data x1 x2 e, corr(corr) sd(sds) n(500)
(obs 500)
gen y = 3*x1 + 2*x2 + e
. corr y x1 x2
(obs=500)
                                x1
                                           x2
```

	+			
У		1.0000		
x1		0.7960	1.0000	
x2		0.6965	0.5000	1.0000

. corr y x1 x2, cov (obs=500)

		У	x1	x2
У	1	404		
x1		64	16	
x2	1	56	8	16

- . * Correct regression
- . reg y x1 x2

Source	SS	df		MS		Number of obs F(2, 497)		500 755.44
Model Residual Total	151696 49899.9993	497		02413		Prob > F R-squared Adj R-squared Root MSE	= = =	0.0000 0.7525
у						[95% Conf.	Int	 erval]
x1 x2 _cons	3 2	.1294 .1294 .4481	885 885	23.17 15.45 -0.00	0.000	2.745588 1.745588 8804284	2.	254412 254412 804284

. * Omitted variable bias

. reg y x1

	SS +	df	MS		Number of obs F(1, 498)		
Model Residual	127744	1 498	127744 148.297187		Prob > F R-squared Adj R-squared	= =	0.0000 0.6337
	201595.999				Root MSE		12.178
у	 Coef.				[95% Conf.	In	terval]
x1 _cons	4	.13628	76 29.35	0.000	3.732231 -1.070006	_	.267769

We see that, when x2 is omitted from the model, the effect of x1 is over-estimated in this case. (In other situations it could be under-estimated). To confirm that Stata got it right,

$$b_1^* = b_1 + b_2 \frac{\hat{C}ov(X_1, X_2)}{\hat{V}(X_1)} = 3 + 2\frac{8}{16} = 4$$

Example 2. Here is an example of a special case where omitting a variable does NOT result in omitted variable bias. I construct a data set similar to what we had before, except x1 and x2 are uncorrelated.

- . * Correct regression
- . reg y x1 x2

Source	SS	df	MS		Number of obs F(2, 497)		500 516.88
Model Residual + Total			51896.0002 100.402413 308		Prob > F R-squared Adj R-squared Root MSE	= = =	0.0000 0.6753 0.6740 10.02
у	Coef.	Std. I	Err. t	P> t	[95% Conf.	In	terval]
x1 x2 _cons	3 2 -4.71e-08	.11214	403 17.83	0.000	2.779672 1.779672 8804285	2	.220328 .220328 8804284

. * X2 omitted but no bias in this case

	SS	df	MS		Number of obs F(1, 498)		500
Model		1 7 498 1	1856.0006 64.329316		Prob > F R-squared Adj R-squared	= =	0.0000 0.4675
Total	!	499	308		Root MSE		12.819
у	Coef.	Std. Er	r. t	P> t	[95% Conf.	In	terval]
x1 _cons	3.71e-08	.143465			2.718128 -1.12636		.281872 1.12636

Inclusion of extraneous variables. Suppose that the "correct" model is

$$y = \alpha + \beta_1 X_1 + \varepsilon$$

If we estimate

$$y = \alpha + b_1 X_1 + e$$

we know that $E(b_1) = \beta_1$, i.e. the regression coefficients is an unbiased estimators of the population parameter.

Suppose, however, the researcher mistakenly believes

$$y = \alpha^* + \beta^*_{1}X_{1} + \beta^*_{2}X_{2} + \varepsilon^*$$

and therefore estimates

$$y = a^* + b^*_1 X_1 + b_2^* X_2 + e^*$$

i.e. X2 is mistakenly added to the model. How does b_1 (the regression estimate from the correctly specified model) compare to b_1^* (the regression estimate from the mis-specified model)? What is $E(b_1^*)$? Is it a biased or unbiased estimator of β_1 ? If biased, how is it biased?

Here is an informal proof: We can think of the "correct" model as being a special case of the "incorrect" model, where $\beta_2 = 0$. It will therefore be the case that $E(b_1^*) = \beta_1$, and $E(b_2^*) = 0$. Hence, addition of extraneous variables does not lead to biased coefficients.

However, adding extraneous (or "junk") variables to the model will result in inflated standard errors and all the problems they create. Recall that, in the two IV case,

$$s_{b_k} = \sqrt{\frac{1 - R_{Y12}^2}{(1 - R_{12}^2) * (N - K - 1)}} * \frac{s_y}{s_x}$$

As the formula suggests, adding irrelevant variables will tend not to increase the numerator, because irrelevant variables will not substantially increase R². However, irrelevant variables will

tend to increase the denominator. The tolerance will be smaller $(1 - R^2_{12})$ and N-K-1 will be smaller.

Example 3. This is similar to the first example, except that x2 has no effect on y.

```
. * Extraneous variables
. clear all
. matrix input corr = (1,.5,0.5,1,0.0,0,1)
. matrix input sds = (4\10)
. corr2data x1 x2 e, corr(corr) sd(sds) n(500)
(obs 500)
. gen y = 3*x1 + e
. corr y x1 x2
(obs=500)
                  x1
              У
_____
       y | 1.0000
      x1 | 0.7682 1.0000
      x2 | 0.3841 0.5000 1.0000
. * Correct regression
. reg y x1
                               Number of obs =
   Source | SS df MS
                                   F( 1, 498) = 717.12
Prob > F = 0.0000
R-squared = 0.5902
  Model | 71856.0006 1 71856.0006
Residual | 49899.9991 498 100.200801
                                      Adj R-squared = 0.5893
_____
    Total | 121756 499 243.999999
                                      Root MSE
      y | Coef. Std. Err. t P>|t| [95% Conf. Interval]
    _____
. * Extraneous variable added
. reg y x1 x2
             SS df MS
                                     Number of obs =
   Source |
                                     F( 2, 497) = 357.84

Prob > F = 0.0000

R-squared = 0.5902
 Model | 71856.0006 2 35928.0003
Residual | 49899.9991 497 100.402413
                                      Adj R-squared = 0.5885
    Total |
            121756 499 243.999999
                                      Root MSE
      y | Coef. Std. Err. t P>|t| [95% Conf. Interval]
_____
```

As you can see the coefficient for x1 did not change but the standard error increased and the t value went down.

Appendix: Another example of omitted variable bias

EXAMPLE: Consider our income/education/job experience example:

. use https://academicweb.nd.edu/~rwilliam/statafiles/reg01.dta, clear
. corr educ jobexp income, cov
(obs=20)

		educ	jobexp	income
		20.05		
jobexp		-2.61316	29.8184	
income		37.0676	14.3108	95.8119

. reg income educ jobexp

Source	SS	df	MS		Number of obs	
Model Residual	1538.22521 282.200265		9.112605 .6000156		F(2, 17) Prob > F R-squared Adj R-squared	= 0.0000 = 0.8450
Total	1820.42548	19 95	.8118671		Root MSE	= 4.0743
income	Coef.	Std. Err	. t	P> t	[95% Conf.	Interval]
educ jobexp _cons	1.933393 .6493654 -7.096855	.2099494 .1721589 3.626412	3.7	7 0.002	1.490438 .2861417 -14.74792	2.376347 1.012589 .5542052

Note that, when both EDUC and JOBEXP are in the equation, $b_1 = 1.933393$, $b_2 = .649365$, Cov(Educ, Jobexp) = -.2613, V(Educ) = 20.05, V(Jobexp) = 29.818. Hence, if we omit Jobexp from the model, the new coefficient b_1^* is

$$b_1^* = b_1 + b_2 \frac{\hat{C}ov(X_1, X_2)}{\hat{V}(X_1)} = 1.933393 + .649365 \frac{-2.613}{20.050} = 1.848765$$

Stata confirms that this is correct:

. reg income educ

Source	SS	df	MS		Number of obs		20
Model	1302.05369 518.371789	1 18	1302.05369		F(1, 18) Prob > F R-squared Adj R-squared	=	45.21 0.0000 0.7152 0.6994
Total	1820.42548	19	95.8118671		Root MSE		5.3664
income	Coef.		 rr. t		[95% Conf.	Int	cerval]
educ _cons	1.84876	.27494	79 6.72	0.000	1.271116 -5.265645		.426404 .540537

Or, if we instead omit EDUC from the equation, for b₂* we get

$$b_1^* = b_2 + b_1 \frac{\hat{C}ov(X_1, X_2)}{\hat{V}(X_2)} = .649365 + .1.933393 \frac{-2.613}{29.818} = .479928616$$

Stata again confirms this:

. reg income jobexp

Source	SS	df	MS		Number of obs F(1, 18)		20 1.39
Model Residual		1	130.495675 93.8849889		Prob > F R-squared	=	0.2538 0.0717
Total	1820.42548	19	95.8118671		Adj R-squared Root MSE		0.0201 9.6894
income	Coef.				[95% Conf.	In	terval]
jobexp _cons	.4799311	.40707	792 1.18	0.254	3753106 6.606476		.335173

If we assume that the model with both EDUC and JOBEXP is correct, omitting one or the other results in the effects of the remaining variable being mis-estimated.

In more complicated models with omitted variables, it will continue to be the case that observed effects represent a confounding of the actual effect with other sources of association.