

Title

survey — Introduction to survey commands

Description

The *Survey Data Reference Manual* organizes the commands alphabetically, making it easy to find individual command entries if you know the name of the command. This overview organizes and presents the commands conceptually, that is, according to the similarities in the functions that they perform.

The following list of commands may have been updated since the release of Stata 9. For an updated list, type the following in an up-to-date Stata:

```
. help survey
```

Survey design tools

svyset	Declare survey design for dataset
svydes	Describe survey data

Descriptive statistics

svy: mean	Estimation of population and subpopulation means
svy: proportion	Estimation of population and subpopulation proportions
svy: ratio	Estimation of population and subpopulation ratios
svy: total	Estimation of population and subpopulation totals
svy: tabulate oneway	One-way tables for survey data
svy: tabulate twoway	Two-way tables for survey data

Regression models

svy: regress	Linear regression for survey data
svy: ivreg	Instrumental variables regression for survey data
svy: intreg	Interval regression for survey data
svy: logistic	Logistic regression, reporting odds ratios, for survey data
svy: logit	Logistic regression, reporting coefficients, for survey data
svy: probit	Probit regression for survey data
svy: mlogit	Multinomial logistic regression for survey data
svy: ologit	Ordered logistic regression for survey data
svy: oprobit	Ordered probit regression for survey data
svy: poisson	Poisson regression for survey data
svy: nbreg	Negative binomial regression for survey data
svy: gnbreg	Generalized negative binomial regression for survey data
svy: heckman	Heckman selection model for survey data
svy: heckprob	Probit regression with selection for survey data

Survey data analysis tools

svy	Overview of the svy prefix command
svy brr	Balanced repeated replication for survey data
svy jackknife	Jackknife estimation for survey data
svy postestimation	Overview of postestimation commands for survey data analysis
estat	Postestimation statistics for survey data
ml for svy	Maximum pseudolikelihood estimation for survey data
svymarkout	Mark observations for exclusion based on survey characteristics

Survey data concepts

variance estimation	Variance estimation for survey data
subpopulation estimation	Subpopulation estimation for survey data
direct standardization	Direct standardization of means, proportions, and ratios
poststratification	Poststratification for survey data

Remarks

Remarks are presented under the headings

Overview
Survey design tools
Descriptive statistics
Regression models
Survey data analysis tools
Survey data concepts

Overview

Stata's facilities for survey data are centered around the **svy** prefix command. Once the design characteristics of a survey dataset are identified with the **svyset** command, the **svy** prefix can be used with supported estimation commands in essentially the same way as the corresponding command for nonsurvey data. For example, where you would normally use the **regress** command to fit a linear regression using nonsurvey data, use **svy: regress** for your survey data.

Why should you use the **svy** prefix command rather than, say, the **mean** command for means or **regress** for linear regression? To answer this question, we need to discuss some of the characteristics of survey design and survey data collection because these characteristics affect how we must perform our analysis if we want to "get it right".

Survey data are characterized by the following:

1. sampling weights, also called probability weights—**pweights** in Stata's syntax
2. cluster sampling
3. stratification

These factors arise from the design of the data collection procedure. Here's a brief description of how these design features affect the analysis of the data:

1. *Sampling weights.* In sample surveys, observations are selected through a random process, but different observations may have different probabilities of selection. Weights are equal to (or proportional to) the inverse of the probability of being sampled. Various postsampling

adjustments to the weights are sometimes made, as well. A weight of w_j for the j th observation means, roughly speaking, that the j th observation represents w_j elements in the population from which the sample was drawn.

Omitting weights from the analysis results in estimates that may be biased, sometimes seriously so. Sampling weights also need to be taken in account when estimating standard errors, and for purposes of testing and inference.

2. *Clustering*. Individuals are not sampled independently in almost all survey designs. Collections of individuals (for example, counties, city blocks, or households) are typically sampled as a group, known as a *cluster*.

There may also be further subsampling within the clusters. For example, counties may be sampled, then city blocks within counties, then households within city blocks, and then finally persons within households. The clusters at the first level of sampling are called *primary sampling units* (PSUs)—in this example, counties are the PSUs. In the absence of clustering, the PSUs are defined to be the individuals or, equivalently, clusters each of size one.

Sampling by cluster implies a sample-to-sample variability of the resulting estimator that is usually greater than that obtained through sampling individually, and this variability must be accounted for when estimating standard errors, testing, or performing other inference.

3. *Stratification*. In surveys, different groups of clusters are often sampled separately. These groups are called *strata*. For example, the 254 counties of a state might be divided into two strata, say, urban counties and rural counties. Then ten counties might be sampled from the urban stratum, and fifteen from the rural stratum.

Sampling is done independently across strata; the stratum divisions are fixed in advance. Thus strata are statistically independent and can be analyzed as such. When the individual strata are more homogenous than the population as a whole, the homogeneity can be exploited to produce smaller (and honestly so) estimates of standard errors.

To put it succinctly: it is important to use sampling weights in order to get the point estimates right. We must consider the weighting, clustering, and stratification of the survey design to get the standard errors right. If our analysis ignores the clustering in our design, we would likely produce standard errors that are smaller than they should be. Stratification can be used to get smaller standard errors for a given total amount of data.

▷ Example 1: A preview of survey data analysis with Stata

We have (fictional) data on American high school seniors (12th graders), and the data were collected according to the following multistage design. In the first stage, counties were independently selected within each state. In the second stage, schools were selected within each chosen county. Within each chosen school, a questionnaire was filled out by every attending high school senior. We've entered all the information into a Stata dataset called `multistage.dta`. The survey design variables are as follows:

1. `state` contains the stratum identifiers
2. `county` contains the first-stage sampling units
3. `ncounties` contains the total number of counties within each state
4. `school` contains the second-stage sampling units
5. `nschools` contains the total number of schools within each county
6. `sampwgt` contains the sampling weight for each sampled individual

Here we load `multistage.dta` into memory and use `svyset` with the above variables to declare that this data is survey data.

```
. use http://www.stata-press.com/data/r9/multistage
. svyset county [pw=sampwt], strata(state) fpc(ncounties) || school, fpc(nschools)
      pweight: sampwt
      VCE: linearized
      Strata 1: state
      SU 1: county
      FPC 1: ncounties
      Strata 2: <one>
      SU 2: school
      FPC 2: nschools
```

Now that the data are `svyset`, we can use the `svy` estimation commands to perform our analysis. In the following, we estimate the mean of weight (in lbs.) for each subpopulation identified by the categories of the `sex` variable (male and female).

```
. svy: mean weight, over(sex)
(running mean on estimation sample)

Survey: Mean estimation
Number of strata =      50      Number of obs   =    4071
Number of PSUs   =    100      Population size = 8.0e+06
                        Design df      =      50

      male: sex = male
      female: sex = female
```

Over	Linearized			
	Mean	Std. Err.	[95% Conf. Interval]	
weight				
male	175.4809	1.116802	173.2377	177.7241
female	146.204	.9004157	144.3955	148.0125

Based on the above results, we are 95% confident that the average weight of male high school seniors is between 173.2 and 177.7 pounds.

Here we use the `test` command to test the hypothesis that the average male is 30 pounds heavier than the average female; however, based on the results we cannot reject this hypothesis at the 5% level.

```
. test [weight]male - [weight]female = 30
Adjusted Wald test
( 1) [weight]male - [weight]female = 30
      F( 1, 50) =    0.23
      Prob > F =    0.6353
```

Survey design tools

Before using `svy`, first take a quick look at [SVY] `svyset`. Use the `svyset` command to specify the variables that identify the survey design characteristics and default method for standard error estimation. Once set, `svy` will automatically use these design specifications until they are cleared or changed, or a new dataset is loaded into memory.

The `svydes` command describes the survey design and is useful in, among other things, tracking down strata with only one sampling unit.

Descriptive statistics

`svy: mean`, `svy: ratio`, `svy: proportion`, and `svy: total` produce estimates of finite-population means, ratios, proportions and totals. `svy: mean`, `svy: ratio`, and `svy: proportion` can also estimate standardized means, ratios, and proportions. Estimates for multiple subpopulations can be obtained using the `over()` option.

`svy: tabulate` can be used to produce one-way and two-way tables with survey data and can also produce tests of independence for two-way contingency tables.

Regression models

Many commands in Stata are used to fit regression models to data, for example `regress` for linear regression, `poisson` for Poisson regression, `logistic` for logistic regression, etc. A subset of these *estimation commands* are supported by `svy`, that is, they may be prefixed by `svy:` in order to produce results appropriate for complex survey data. Whereas `poisson` is used with standard, nonsurvey data, `svy: poisson` is used with survey data. In what follows we refer to any estimation command unprefixed by `svy:` as the standard command. A standard command prefixed by `svy:` is referred to as a `svy` command.

Most standard commands (and all standard commands supported by `svy`) allow `pweights` and the `cluster(varname)` option, where `varname` corresponds to the `psu` variable that you `svyset`. If your survey data exhibit only sampling weights and/or first-stage clusters, you can get by with using the standard command with `pweights` and/or `cluster()`. Your parameter estimates will always be identical to those you would have obtained from the `svy` command, and the standard command uses the same robust (linearization) variance estimator as the `svy` command with a similarly `svyset` design.

Most standard commands are also fit using maximum-likelihood methodology. When used with independently distributed, nonweighted data, the likelihood to be maximized is reflective of the joint probability distribution of the data given the chosen model. With complex survey data, however, this interpretation of the likelihood is no longer valid, as survey data are either weighted, not independently distributed, or both. With survey data, (valid) parameter estimates are obtained using the independence-assuming likelihood and weighting if necessary. Since the probabilistic interpretation no longer holds, the likelihood here is instead called a *pseudolikelihood*. See Skinner (1989, section 3.4.4) for a discussion of maximum pseudolikelihood estimators.

Below we highlight the other features of `svy` commands.

1. `svy` commands handle stratified sampling, but none of the standard commands do. Since stratification usually makes standard errors smaller, ignoring stratification is usually conservative. So, not using `svy` with stratified sample data is not a terrible thing to do. However, to get the smallest possible “honest” standard-error estimates for stratified sampling, use `svy`.

2. `svy` commands use t statistics with $n - L$ degrees of freedom to test the significance of coefficients, where n is the total number of sampled PSUs (clusters) and L is the number of strata in the first stage. Some of the standard commands use t statistics, but most use z statistics. If the standard command uses z statistics for its standard variance estimator, then it also uses z statistics with the robust (linearization) variance estimator. Strictly speaking, t statistics are appropriate with the robust (linearization) variance estimator; see [P] `_robust` for the theoretical rationale. But, using z rather than t statistics only yields a nontrivial difference when there is a small number of clusters (< 50). If a regression model command uses t statistics and the `cluster()` option is specified, then the degrees of freedom used are the same as that of the `svy` command (in the absence of stratification).
3. `svy` commands produce an adjusted Wald test for the model test, and `test` can be used to produce adjusted Wald tests for other hypotheses after `svy` commands. Only unadjusted Wald tests are available if the `svy` prefix is not used. The adjustment can be important when the degrees of freedom $n - L$ are small relative to the dimension of the test. (If the dimension is one, then the adjusted and unadjusted Wald tests are identical.) This fact along with point 2 make it important to use the `svy` command if the number of sampled PSUs (clusters) is small (< 50).
4. `svy: regress` differs slightly from `regress` and `svy: ivreg` differs slightly from `ivreg` in that they use different multipliers for the variance estimator. `regress` and `ivreg` use a multiplier of $\{(N - 1)/(N - k)\}\{n/(n - 1)\}$, where N is the number of observations, n is the number of clusters (PSUs), and k is the number of regressors including the constant. `svy: regress` and `svy: ivreg` use $n/(n - 1)$ instead. Thus they produce slightly different standard errors. The $(N - 1)/(N - k)$ is ad hoc and has no rigorous theoretical justification; hence, the purist `svy` commands do not use it. The `svy` commands tacitly assume that $N \gg k$. If $(N - 1)/(N - k)$ is not close to 1, you may be well advised to use `regress` or `ivreg` so that some punishment is inflicted on your variance estimates. Note that maximum likelihood estimators in Stata (e.g., `logit`) do no such adjustment, but rely on the sensibilities of the analyst to ensure that N is reasonably larger than k . Thus the maximum pseudolikelihood estimators (e.g., `svy: logit`) produce exactly the same standard errors as the corresponding maximum likelihood commands (e.g., `logit`), but p -values are slightly different because of point 2.
5. `svy` commands can produce proper estimates for subpopulations through use of the `subpop()` option. Use of an `if` restriction with `svy` or standard commands can yield incorrect standard error estimates for subpopulations. Often an `if` restriction will yield exactly the same standard error as `subpop()`; most other times, the two standard errors will be slightly different; but, in some cases—usually for thinly sampled subpopulations—the standard errors can be appreciably different. Hence, the `svy` command with the `subpop()` option should be used to obtain estimates for thinly sampled subpopulations. See [SVY] **subpopulation estimation** for more information.
6. `svy` commands handle zero sampling weights properly. Standard commands ignore any observation with a weight of zero. Usually, this will yield exactly the same standard errors, but sometimes they will differ. Sampling weights of zero can arise from various postsampling adjustment procedures. If the sum of weights for one or more PSUs is zero, `svy` and standard commands will produce different standard errors, but usually this difference is very small.
7. You can `svyset iweights` and let these weights be negative. Negative sampling weights can arise from various postsampling adjustment procedures. If you want to use negative sampling weights, then you must `svyset iweights` instead of `pweights`; no standard command will allow negative sampling weights.
8. The `svy` commands compute finite population corrections (FPC).

9. After a `svy` command, `estat effects` will compute the design effects `DEFF` and `DEFT` and the misspecification effects `MEFF` and `MEFT`.
10. `svy` commands can perform variance estimation that accounts for multiple stages of clustered sampling.
11. `svy` commands can perform variance estimation that accounts for poststratification adjustments to the sampling weights.

Survey data analysis tools

Stata's suite of survey-data commands is governed by the `svy` prefix command. `svy` runs the supplied estimation command while accounting for the survey design characteristics in the point estimates and the variance estimator. The three available variance estimation methods are balanced repeated replication (BRR), the jackknife, and first-order Taylor linearization. By default, `svy` computes standard errors using the linearized variance estimator—so called because it is based on a first-order Taylor series linear approximation. In the nonsurvey context, we refer to this variance estimator as the *robust* variance estimator, otherwise known in Stata as the Huber/White/sandwich estimator; see [P] `_robust`.

The `svy brr` and `svy jackknife` prefix commands can be used with those commands that may not be fully supported by `svy` but are compatible with BRR and the jackknife; see [SVY] `svy brr` and [SVY] `svy jackknife`.

All the standard postestimation commands (e.g., `estat`, `lincom`, `nlcom`, `test`, `testnl`) are also available after `svy`.

`estat` has specific subroutines for use after `svy`. `estat svyset` reports the survey design settings used to produce the current estimation results. `estat effects` and `estat lceffects` report a table of design and misspecification effects for point estimates and linear combinations of point estimates, respectively. `estat size` reports a table of sample and subpopulation sizes after `svy: mean`, `svy: proportion`, `svy: ratio`, and `svy: total`.

The `ml` command can be used to fit a pseudolikelihood model. When maximum pseudolikelihood is carried out using `ml`, the weighting during estimation and postestimation linearization is performed automatically, provided that the user specifies the appropriate survey options to `ml`; see [R] `ml` for details.

`svymarkout` is a programmer's command that resets the values in a variable that identifies the estimation sample, dropping observations for which any of the survey-characteristic variables contain missing values. This tool is most helpful for developing estimation commands that use `ml` to fit models using maximum pseudolikelihood.

Survey data concepts

The variance estimation methods used by Stata are discussed in [SVY] **variance estimation**.

See [SVY] **subpopulation estimation** for an explanation of why you should use the `subpop()` option instead of the `if` and `in` options.

The weight adjusting methods for direct standardization and poststratification are discussed in [SVY] **direct standardization** and [SVY] **poststratification**.

For more detailed introductions to complex survey data analysis, see Scheaffer, Mendenhall, and Ott (1996), Stuart (1984), Williams (1978), and Levy and Lemeshow (1999). Advanced treatments and discussion of important special topics are given by Cochran (1977), Korn and Graubard (1999).

Särndal, Swensson, and Wretman (1992), Shao and Tu (1995), Skinner, Holt, and Smith (1989), Thompson (2002), and Wolter (1985).

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William Gemmell Cochran (1909–1980) was born in Rutherglen, Scotland, and educated at the Universities of Glasgow and Cambridge. He accepted a post at Rothamsted before finishing his doctorate. Cochran emigrated to the United States in 1939 and worked at Iowa State, North Carolina State, Johns Hopkins, and Harvard. He made many major contributions across several fields of statistics, including experimental design, the analysis of counted data, sample surveys and observational studies, and was author or co-author (with Gertrude M. Cox and George W. Snedecor) of various widely used texts.

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Also See

- Complementary:** [R] estat, [R] jackknife, [R] lincom, [R] ml, [R] nlcom,
 [R] predict, [R] predictnl, [R] test, [R] testnl
- Background:** [P] _robust